

Probabilistic classification

CE-717: Machine Learning
Sharif University of Technology

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Topics

- ▶ Probabilistic approach
 - ▶ Bayes decision theory
 - ▶ Generative models
 - ▶ Gaussian Bayes classifier
 - ▶ Naïve Bayes
 - ▶ Discriminative models
 - ▶ Logistic regression

Classification problem: probabilistic view

- ▶ Each feature as a random variable
- ▶ Class label also as a random variable
- ▶ We observe the feature values for a random sample and we intend to find its class label
 - ▶ Evidence: feature vector x
 - ▶ Query: class label

Definitions

- ▶ Posterior probability: $p(\mathcal{C}_k | \mathbf{x})$
- ▶ Likelihood or class conditional probability: $p(\mathbf{x} | \mathcal{C}_k)$
- ▶ Prior probability: $p(\mathcal{C}_k)$

$p(\mathbf{x})$: pdf of feature vector \mathbf{x} ($p(\mathbf{x}) = \sum_{k=1}^K p(\mathbf{x} | \mathcal{C}_k) p(\mathcal{C}_k)$)

$p(\mathbf{x} | \mathcal{C}_k)$: pdf of feature vector \mathbf{x} for samples of class \mathcal{C}_k

$p(\mathcal{C}_k)$: probability of the label be \mathcal{C}_k

Bayes decision rule

$K = 2$

If $P(\mathcal{C}_1|\mathbf{x}) > P(\mathcal{C}_2|\mathbf{x})$ decide \mathcal{C}_1
otherwise decide \mathcal{C}_2

$$p(error|\mathbf{x}) = \begin{cases} p(\mathcal{C}_2|\mathbf{x}) & \text{if we decide } \mathcal{C}_1 \\ P(\mathcal{C}_1|\mathbf{x}) & \text{if we decide } \mathcal{C}_2 \end{cases}$$

- If we use Bayes decision rule:

$$P(error|\mathbf{x}) = \min\{P(\mathcal{C}_1|\mathbf{x}), P(\mathcal{C}_2|\mathbf{x})\}$$

Using Bayes rule, for each \mathbf{x} , $P(error|\mathbf{x})$ is as small as possible and thus this rule minimizes the probability of error

Optimal classifier

- ▶ The optimal decision is the one that minimizes the expected number of mistakes
- ▶ We show that Bayes classifier is an optimal classifier

Bayes decision rule

Minimizing misclassification rate

► Decision regions: $\mathcal{R}_k = \{\mathbf{x} | \alpha(\mathbf{x}) = k\}$

$K = 2$

► All points in \mathcal{R}_k are assigned to class \mathcal{C}_k

$$\begin{aligned} p(\text{error}) &= E_{\mathbf{x}, y}[I(\alpha(\mathbf{x}) \neq y)] \\ &= p(\mathbf{x} \in \mathcal{R}_1, \mathcal{C}_2) + p(\mathbf{x} \in \mathcal{R}_2, \mathcal{C}_1) \\ &= \int_{\mathcal{R}_1} p(\mathbf{x}, \mathcal{C}_2) d\mathbf{x} + \int_{\mathcal{R}_2} p(\mathbf{x}, \mathcal{C}_1) d\mathbf{x} \\ &= \int_{\mathcal{R}_1} p(\mathcal{C}_2 | \mathbf{x}) p(\mathbf{x}) d\mathbf{x} + \int_{\mathcal{R}_2} p(\mathcal{C}_1 | \mathbf{x}) p(\mathbf{x}) d\mathbf{x} \end{aligned}$$

Choose class with highest $p(\mathcal{C}_k | \mathbf{x})$ as $\alpha(\mathbf{x})$

Bayes minimum error

- ▶ Bayes minimum error classifier:

$$\min_{\alpha(\cdot)} E_{\mathbf{x},y}[I(\alpha(\mathbf{x}) \neq y)] \quad \text{Zero-one loss}$$

- ▶ If we know the probabilities in advance then the above optimization problem will be solved easily.
 - ▶ $\alpha(\mathbf{x}) = \operatorname{argmax}_y p(y|\mathbf{x})$
- ▶ In practice, we can estimate $p(y|\mathbf{x})$ based on a set of training samples \mathcal{D}

Bayes theorem

► Bayes' theorem

$$p(\mathcal{C}_k|\mathbf{x}) = \frac{p(\mathbf{x}|\mathcal{C}_k)p(\mathcal{C}_k)}{p(\mathbf{x})}$$

Diagram illustrating the components of Bayes' theorem:

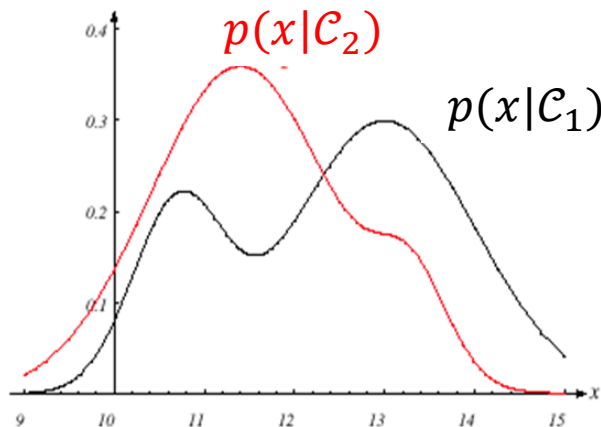
- Posterior: $p(\mathcal{C}_k|\mathbf{x})$
- Likelihood: $p(\mathbf{x}|\mathcal{C}_k)$
- Prior: $p(\mathcal{C}_k)$

- Posterior probability: $p(\mathcal{C}_k|\mathbf{x})$
- Likelihood or class conditional probability: $p(\mathbf{x}|\mathcal{C}_k)$
- Prior probability: $p(\mathcal{C}_k)$

$p(\mathbf{x})$: pdf of feature vector \mathbf{x} ($p(\mathbf{x}) = \sum_{k=1}^K p(\mathbf{x}|\mathcal{C}_k)p(\mathcal{C}_k)$)
 $p(\mathbf{x}|\mathcal{C}_k)$: pdf of feature vector \mathbf{x} for samples of class \mathcal{C}_k
 $p(\mathcal{C}_k)$: probability of the label be \mathcal{C}_k

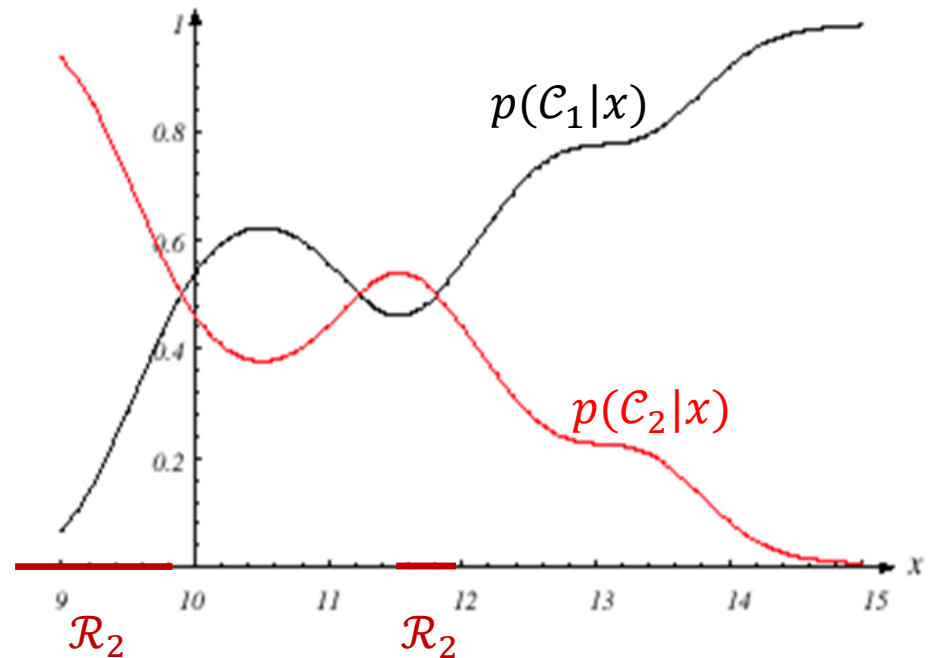
Bayes decision rule: example

- Bayes decision: Choose the class with highest $p(\mathcal{C}_k|x)$



$$p(\mathcal{C}_1) = \frac{2}{3}$$



$$p(\mathcal{C}_2) = \frac{1}{3}$$



$$p(\mathcal{C}_k|x) = \frac{p(x|\mathcal{C}_k)p(\mathcal{C}_k)}{p(x)}$$

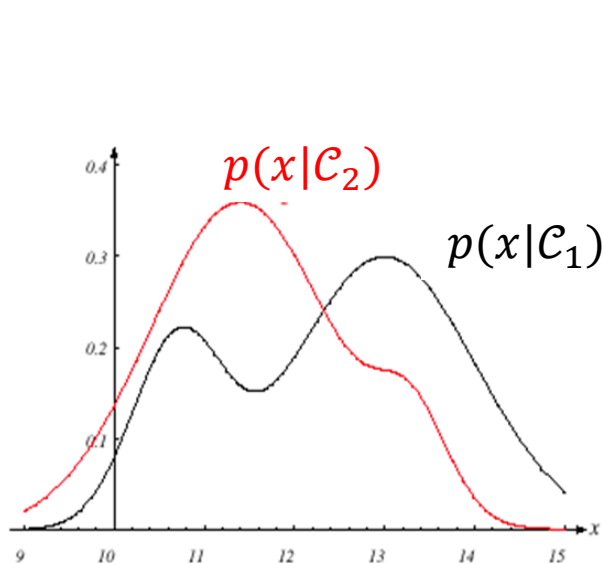
$$p(x) = p(\mathcal{C}_1)p(x|\mathcal{C}_1) + p(\mathcal{C}_2)p(x|\mathcal{C}_2)$$

Bayesian decision rule

- ▶ If $P(\mathcal{C}_1|\mathbf{x}) > P(\mathcal{C}_2|\mathbf{x})$ decide \mathcal{C}_1
otherwise decide \mathcal{C}_2
 Equivalent
- ▶ If $\frac{p(\mathbf{x}|\mathcal{C}_1)P(\mathcal{C}_1)}{p(\mathbf{x})} > \frac{p(\mathbf{x}|\mathcal{C}_2)P(\mathcal{C}_2)}{p(\mathbf{x})}$ decide \mathcal{C}_1
otherwise decide \mathcal{C}_2
 Equivalent
- ▶ If $p(\mathbf{x}|\mathcal{C}_1)P(\mathcal{C}_1) > p(\mathbf{x}|\mathcal{C}_2)P(\mathcal{C}_2)$ decide \mathcal{C}_1
otherwise decide \mathcal{C}_2

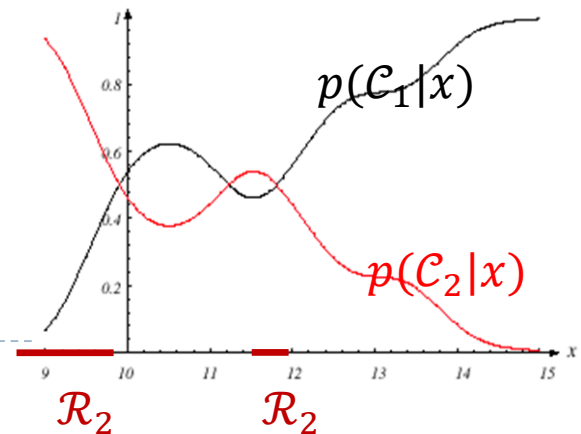
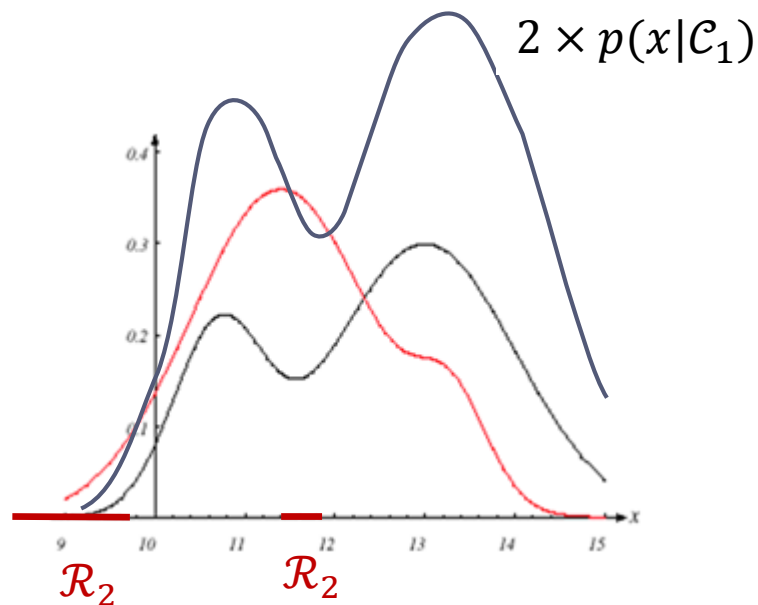
Bayes decision rule: example

- Bayes decision: Choose the class with highest $p(\mathcal{C}_k|x)$



$$p(\mathcal{C}_1) = \frac{2}{3}$$

$$p(\mathcal{C}_2) = \frac{1}{3}$$

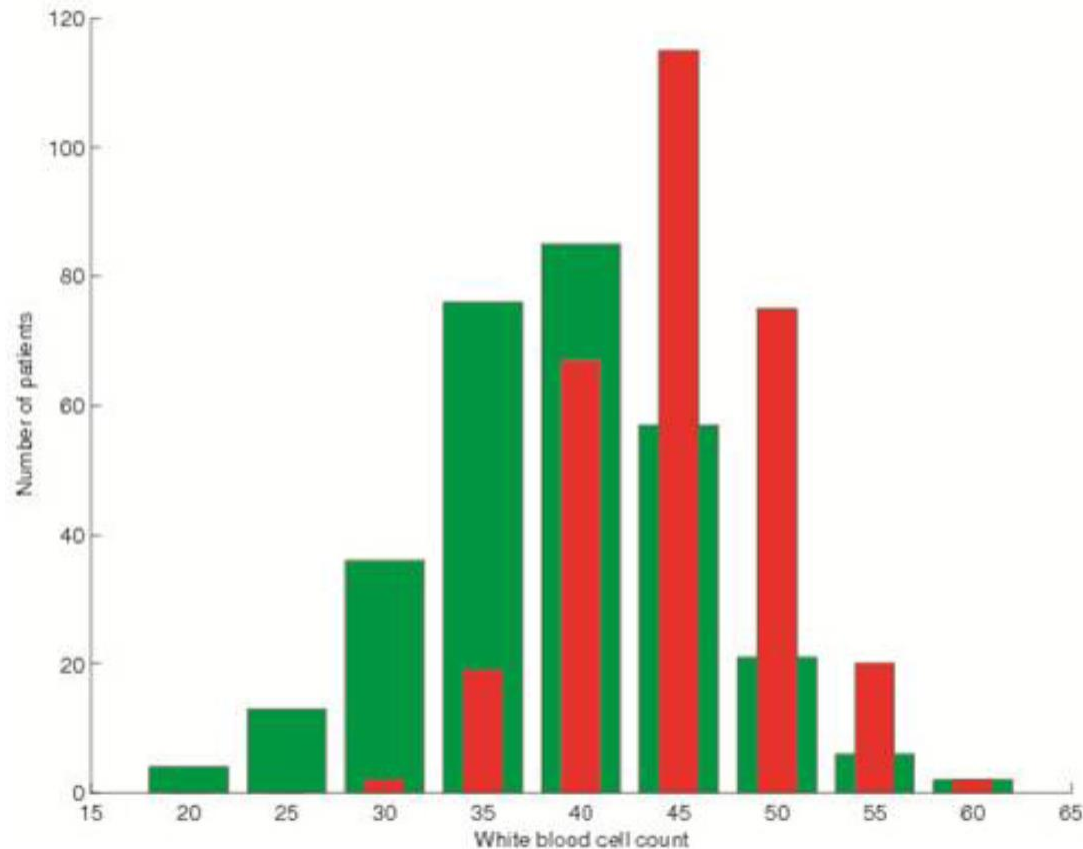


Bayes Classifier

- ▶ Simple Bayes classifier: estimate posterior probability of each class
- ▶ What should the decision criterion be?
 - ▶ Choose class with highest $p(\mathcal{C}_k|x)$
- ▶ The optimal decision is the one that minimizes the expected number of mistakes

Diabetes example

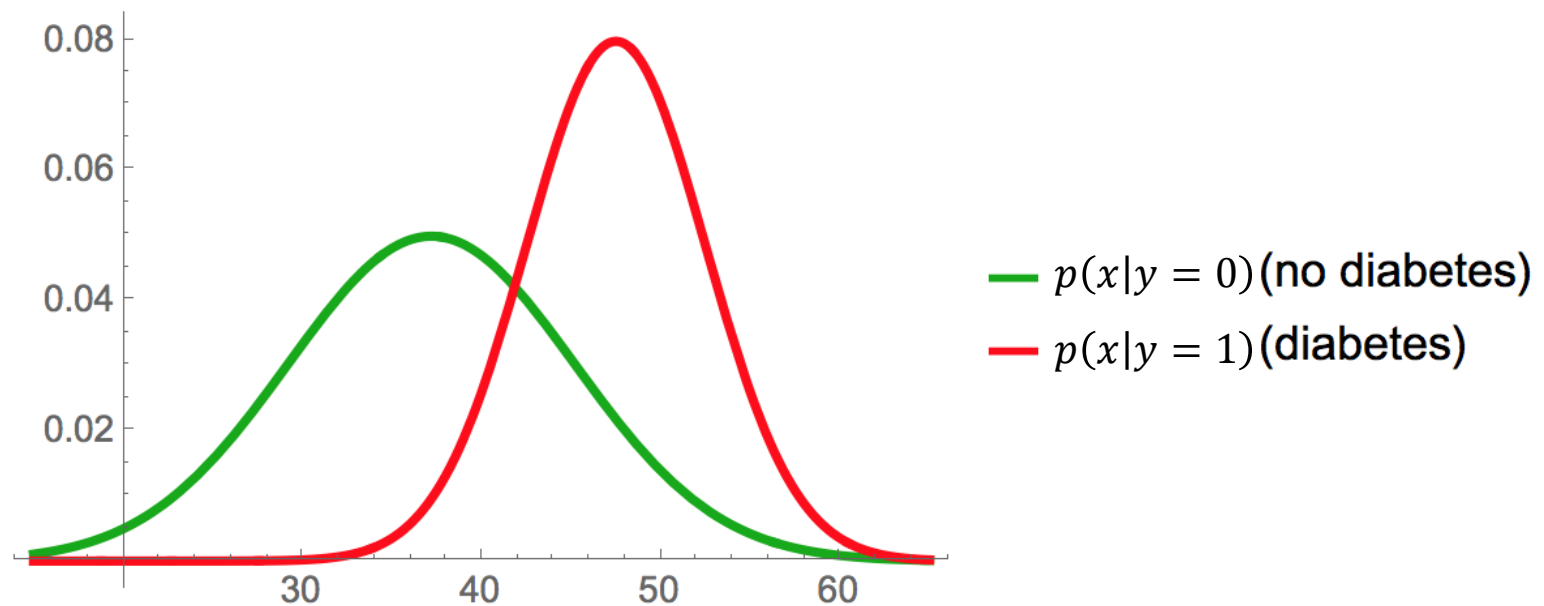
- ▶ white blood cell count



Diabetes example

- ▶ Doctor has a prior $p(y = 1) = 0.2$
 - ▶ Prior: In the absence of any observation, what do I know about the probability of the classes?
- ▶ A patient comes in with white blood cell count x
- ▶ Does the patient have diabetes $p(y = 1|x)$?
 - ▶ given a new observation, we still need to compute the posterior

Diabetes example

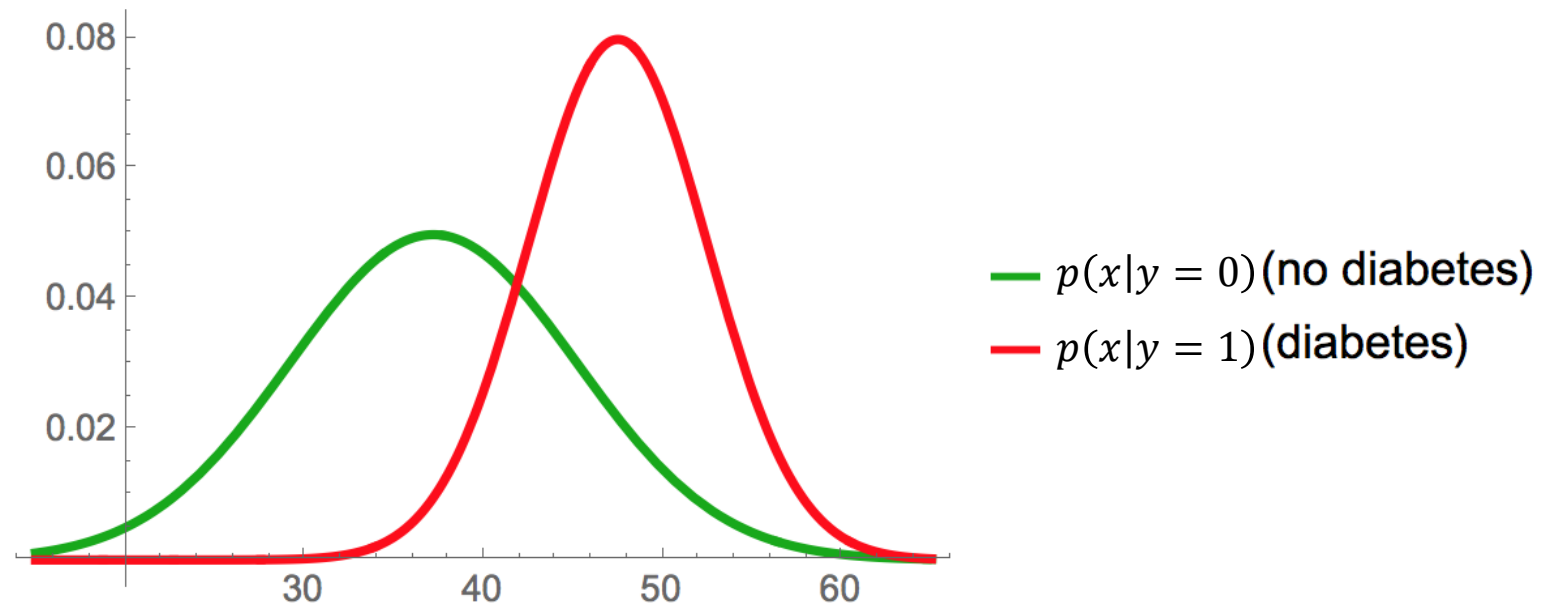


Estimate probability densities from data

- ▶ If we assume Gaussian distributions for $p(x|\mathcal{C}_1)$ and $p(x|\mathcal{C}_2)$
- ▶ Recall that for samples $\{x^{(1)}, \dots, x^{(N)}\}$, if we assume a Gaussian distribution, the MLE estimates will be

$$\begin{aligned}\mu &= \frac{1}{N} \sum_{n=1}^N x^{(n)} \\ \sigma^2 &= \frac{1}{N} \sum_{n=1}^N (x^{(n)} - \mu)^2\end{aligned}$$

Diabetes example



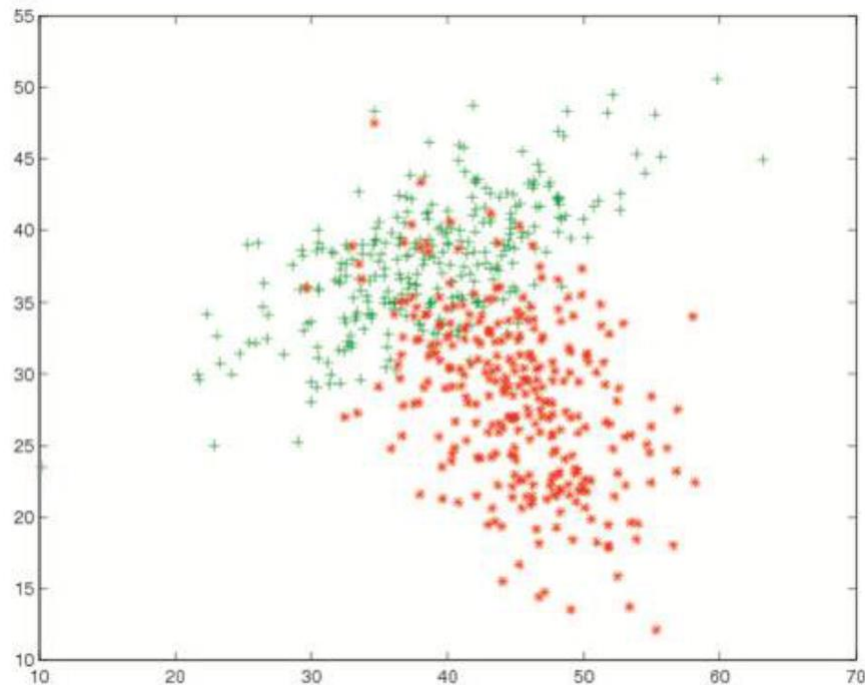
$$p(x|y = 1) = N(\mu_1, \sigma_1^2)$$

$$\mu_1 = \frac{\sum_{n: y^{(n)}=1} x^{(n)}}{\sum_{n: y^{(n)}=1} 1} = \frac{\sum_{n: y^{(n)}=1} x^{(n)}}{N_1}$$

$$\sigma_1^2 = \frac{\sum_{n: y^{(n)}=1} (x^{(n)} - \mu_1)^2}{N_1}$$

Diabetes example

- ▶ Add a second observation: Plasma glucose value



Generative approach for this example

- ▶ Multivariate Gaussian distributions for $p(\mathbf{x}|\mathcal{C}_k)$:

$$p(\mathbf{x}|y = k) = \frac{1}{(2\pi)^{d/2} |\Sigma|^{1/2}} \exp\left\{-\frac{1}{2} (\mathbf{x} - \boldsymbol{\mu}_k)^T \boldsymbol{\Sigma}_k^{-1} (\mathbf{x} - \boldsymbol{\mu}_k)\right\}$$

$$k = 1, 2$$

- ▶ Prior distribution $p(\mathbf{x}|\mathcal{C}_k)$:

- ▶ $p(y = 1) = \pi, \quad p(y = 0) = 1 - \pi$

MLE for multivariate Gaussian

- ▶ For samples $\{x^{(1)}, \dots, x^{(N)}\}$, if we assume a multivariate Gaussian distribution, the MLE estimates will be:

$$\boldsymbol{\mu} = \frac{\sum_{n=1}^N \mathbf{x}^{(n)}}{N}$$

$$\boldsymbol{\Sigma} = \frac{1}{N} \sum_{n=1}^N (\mathbf{x}^{(n)} - \boldsymbol{\mu})(\mathbf{x}^{(n)} - \boldsymbol{\mu})^T$$

Generative approach: example

Maximum likelihood estimation ($D = \{(\mathbf{x}^{(n)}, y^{(n)})\}_{n=1}^N$):

▶ $\pi = \frac{N_1}{N}$

▶ $\boldsymbol{\mu}_1 = \frac{\sum_{n=1}^N y^{(n)} \mathbf{x}^{(n)}}{N_1}, \boldsymbol{\mu}_2 = \frac{\sum_{n=1}^N (1 - y^{(n)}) \mathbf{x}^{(n)}}{N_2}$

▶ $\boldsymbol{\Sigma}_1 = \frac{1}{N_1} \sum_{n=1}^N y^{(n)} (\mathbf{x}^{(n)} - \boldsymbol{\mu})(\mathbf{x}^{(n)} - \boldsymbol{\mu})^T$

▶ $\boldsymbol{\Sigma}_2 = \frac{1}{N_2} \sum_{n=1}^N (1 - y^{(n)}) (\mathbf{x}^{(n)} - \boldsymbol{\mu})(\mathbf{x}^{(n)} - \boldsymbol{\mu})^T$

$$N_1 = \sum_{n=1}^N y^{(n)}$$

$$N_2 = N - N_1$$

Decision boundary for Gaussian Bayes classifier

$$p(\mathcal{C}_1|\mathbf{x}) = p(\mathcal{C}_2|\mathbf{x})$$

$$p(\mathcal{C}_k|\mathbf{x}) = \frac{p(\mathbf{x}|\mathcal{C}_k)p(\mathcal{C}_k)}{p(\mathbf{x})}$$

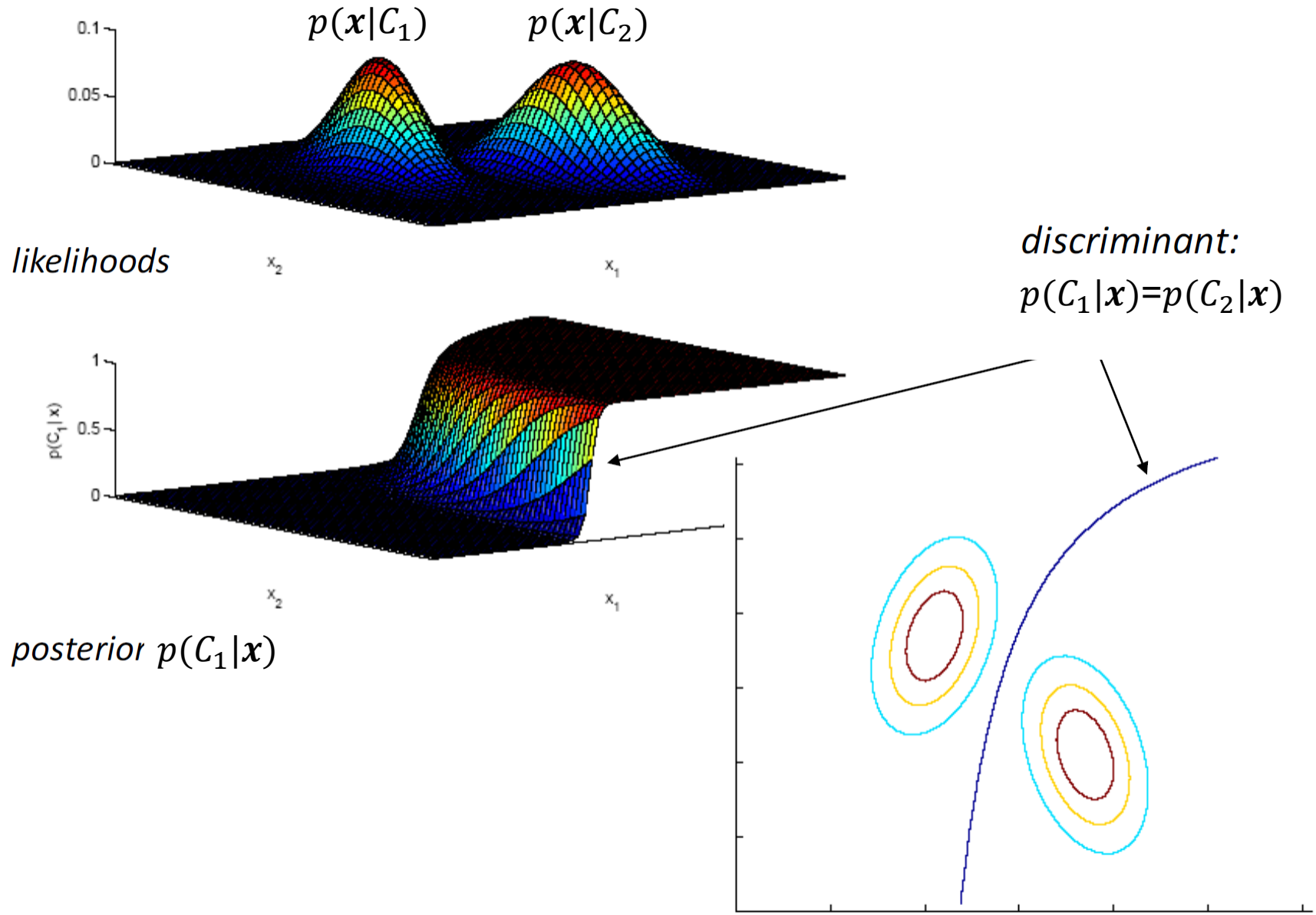
$$\ln p(\mathcal{C}_1|\mathbf{x}) = \ln p(\mathcal{C}_2|\mathbf{x})$$

$$\begin{aligned} \ln p(\mathbf{x}|\mathcal{C}_1) + \ln p(\mathcal{C}_1) - \ln p(\mathbf{x}) \\ = \ln p(\mathbf{x}|\mathcal{C}_2) + \ln p(\mathcal{C}_2) - \ln p(\mathbf{x}) \end{aligned}$$

$$\ln p(\mathbf{x}|\mathcal{C}_1) + \ln p(\mathcal{C}_1) = \ln p(\mathbf{x}|\mathcal{C}_2) + \ln p(\mathcal{C}_2)$$

$$\begin{aligned} \ln p(\mathbf{x}|\mathcal{C}_k) \\ = -\frac{d}{2} \ln 2\pi - \frac{1}{2} \ln |\boldsymbol{\Sigma}_k^{-1}| - \frac{1}{2} (\mathbf{x} - \boldsymbol{\mu}_k)^T \boldsymbol{\Sigma}_k^{-1} (\mathbf{x} - \boldsymbol{\mu}_k) \end{aligned}$$

Decision boundary



Shared covariance matrix

- ▶ When classes share a single covariance matrix $\Sigma = \Sigma_1 = \Sigma_2$

$$p(\mathbf{x}|C_k) = \frac{1}{(2\pi)^{d/2} |\Sigma|^{1/2}} \exp\left\{-\frac{1}{2} (\mathbf{x} - \boldsymbol{\mu}_k)^T \Sigma^{-1} (\mathbf{x} - \boldsymbol{\mu}_k)\right\}$$

$$k = 1, 2$$

- ▶ $p(C_1) = \pi, \quad p(C_2) = 1 - \pi$

Likelihood

$$\begin{aligned} & \prod_{n=1}^N p(\mathbf{x}^{(n)}, y^{(n)} | \pi, \boldsymbol{\mu}_1, \boldsymbol{\mu}_2, \boldsymbol{\Sigma}) \\ &= \prod_{n=1}^N p(\mathbf{x}^{(n)} | y^{(n)}, \boldsymbol{\mu}_1, \boldsymbol{\mu}_2, \boldsymbol{\Sigma}) p(y^{(n)} | \pi) \end{aligned}$$

Shared covariance matrix

- ▶ Maximum likelihood estimation ($D = \{(\mathbf{x}^{(i)}, y^{(i)})\}_{i=1}^n$):

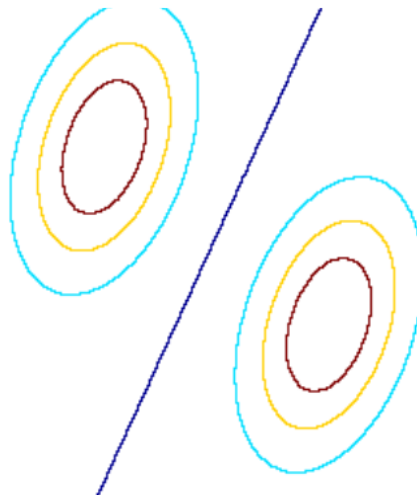
$$\begin{aligned}\pi &= \frac{N_1}{N} \\ \boldsymbol{\mu}_1 &= \frac{\sum_{n=1}^N y^{(n)} \mathbf{x}^{(n)}}{N_1} \\ \boldsymbol{\mu}_2 &= \frac{\sum_{n=1}^N (1 - y^{(n)}) \mathbf{x}^{(n)}}{N_2}\end{aligned}$$

$$\boldsymbol{\Sigma} = \frac{1}{N} \left(\sum_{n \in C_1} (\mathbf{x}^{(n)} - \boldsymbol{\mu}_1)(\mathbf{x}^{(n)} - \boldsymbol{\mu}_1)^T + \sum_{n \in C_2} (\mathbf{x}^{(n)} - \boldsymbol{\mu}_2)(\mathbf{x}^{(n)} - \boldsymbol{\mu}_2)^T \right)$$

Decision boundary when shared covariance matrix

$$\ln p(\mathbf{x}|\mathcal{C}_1) + \ln p(\mathcal{C}_1) = \ln p(\mathbf{x}|\mathcal{C}_2) + \ln p(\mathcal{C}_2)$$

$$\begin{aligned} \ln p(\mathbf{x}|\mathcal{C}_k) \\ = -\frac{d}{2} \ln 2\pi - \frac{1}{2} \ln |\boldsymbol{\Sigma}_k^{-1}| - \frac{1}{2} (\mathbf{x} - \boldsymbol{\mu}_k)^T \boldsymbol{\Sigma}^{-1} (\mathbf{x} - \boldsymbol{\mu}_k) \end{aligned}$$



Bayes decision rule

Multi-class misclassification rate

- ▶ Multi-class problem: Probability of error of Bayesian decision rule
 - ▶ Simpler to compute the probability of correct decision

$$P(\text{error}) = 1 - P(\text{correct})$$

$$\begin{aligned} P(\text{Correct}) &= \sum_{i=1}^K \int_{\mathcal{R}_i} p(\mathbf{x}, \mathcal{C}_i) d\mathbf{x} \\ &= \sum_{i=1}^K \int_{\mathcal{R}_i} p(\mathcal{C}_i | \mathbf{x}) p(\mathbf{x}) d\mathbf{x} \end{aligned}$$

\mathcal{R}_i : the subset of feature space assigned to the class \mathcal{C}_i using the classifier

Bayes minimum error

► Bayes minimum error classifier:

$$\min_{\alpha(\cdot)} E_{\mathbf{x},y}[I(\alpha(\mathbf{x}) \neq y)] \quad \text{Zero-one loss}$$

$$\alpha(\mathbf{x}) = \operatorname{argmax}_y p(y|\mathbf{x})$$

Minimizing Bayes risk (expected loss)

$$\begin{aligned} & E_{\mathbf{x},y}[L(\alpha(\mathbf{x}), y)] \\ &= \int \sum_{j=1}^K L(\alpha(\mathbf{x}), \mathcal{C}_j) p(\mathbf{x}, \mathcal{C}_j) d\mathbf{x} \\ &= \int p(\mathbf{x}) \underbrace{\sum_{j=1}^K L(\alpha(\mathbf{x}), \mathcal{C}_j) p(\mathcal{C}_j | \mathbf{x})}_{\text{conditional risk}} d\mathbf{x} \end{aligned}$$

for each \mathbf{x} minimize it that is called conditional risk

- Bayes minimum loss (risk) decision rule: $\hat{\alpha}(\mathbf{x})$

$$\hat{\alpha}(\mathbf{x}) = \operatorname{argmin}_{i=1,\dots,K} \sum_{j=1}^K \underset{\substack{\downarrow \\ \text{loss of assigning } \mathbf{x} \text{ to } \mathcal{C}_i \text{ when } \mathcal{C}_j \text{ is correct}}}{L_{ij}} p(\mathcal{C}_j | \mathbf{x})$$

The loss of assigning a sample to \mathcal{C}_i where the correct class is \mathcal{C}_j

Minimizing expected loss: special case (loss = misclassification rate)

► Problem definition for this special case:

- If action $\alpha(\mathbf{x}) = i$ is taken and the true category is \mathcal{C}_j , then the decision is correct if $i = j$ and otherwise it is incorrect.

- Zero-one loss function:

$$L_{ij} = 1 - \delta_{ij} = \begin{cases} 0 & i = j \\ 1 & \text{o. w.} \end{cases}$$

$$\hat{\alpha}(\mathbf{x}) = \operatorname{argmin}_{i=1,\dots,K} \sum_{j=1}^K L_{ij} p(\mathcal{C}_j | \mathbf{x})$$

$$= \operatorname{argmin}_{i=1,\dots,K} 0 \times p(\mathcal{C}_i | \mathbf{x}) + \sum_{j \neq i} p(\mathcal{C}_j | \mathbf{x})$$

$$= \operatorname{argmin}_{i=1,\dots,K} 1 - p(\mathcal{C}_i | \mathbf{x}) = \operatorname{argmax}_{i=1,\dots,K} p(\mathcal{C}_i | \mathbf{x})$$

Probabilistic discriminant functions

- ▶ **Discriminant functions:** A popular way of representing a classifier

- ▶ A discriminant function $f_i(\mathbf{x})$ for each class \mathcal{C}_i ($i = 1, \dots, K$):
 - ▶ \mathbf{x} is assigned to class \mathcal{C}_i if:

$$f_i(\mathbf{x}) > f_j(\mathbf{x}) \quad \forall j \neq i$$

- ▶ Representing Bayesian classifier using discriminant functions:

- ▶ Classifier minimizing error rate: $f_i(\mathbf{x}) = P(\mathcal{C}_i|\mathbf{x})$
- ▶ Classifier minimizing risk: $f_i(\mathbf{x}) = -\sum_{j=1}^K L_{ij}p(\mathcal{C}_j|\mathbf{x})$

Naïve Bayes classifier

- ▶ Generative methods
 - ▶ High number of parameters
- ▶ Assumption: Conditional independence

$$p(\mathbf{x}|C_k) = p(x_1|C_k) \times p(x_2|C_k) \times \cdots \times p(x_d|C_k)$$

Naïve Bayes classifier

- ▶ In the decision phase, it finds the label of \mathbf{x} according to:

$$\operatorname{argmax}_{k=1,\dots,K} p(C_k|\mathbf{x})$$
$$\operatorname{argmax}_{k=1,\dots,K} p(C_k) \prod_{i=1}^n p(x_i|C_k)$$

$$p(\mathbf{x}|C_k) = p(x_1|C_k) \times p(x_2|C_k) \times \dots \times p(x_d|C_k)$$
$$p(C_k|\mathbf{x}) \propto p(C_k) \prod_{i=1}^n p(x_i|C_k)$$

Naïve Bayes classifier

- ▶ Finds d univariate distributions $p(x_1|C_k), \dots, p(x_d|C_k)$ instead of finding one multi-variate distribution $p(\mathbf{x}|C_k)$
 - ▶ Example 1: For Gaussian class-conditional density $p(\mathbf{x}|C_k)$, it finds $d + d$ (mean and sigma parameters on different dimensions) instead of $d + \frac{d(d+1)}{2}$ parameters
 - ▶ Example 2: For Bernoulli class-conditional density $p(\mathbf{x}|C_k)$, it finds d (mean parameters on different dimensions) instead of $2^d - 1$ parameters
- ▶ It first estimates the class conditional densities $p(x_1|C_k), \dots, p(x_d|C_k)$ and the prior probability $p(C_k)$ for each class ($k = 1, \dots, K$) based on the training set.

Naïve Bayes: discrete example

▶ $p(H = \text{Yes}) = 0.3$

▶ $p(D = \text{Yes} | H = \text{Yes}) = \frac{1}{3}$

▶ $p(S = \text{Yes} | H = \text{Yes}) = \frac{2}{3}$

▶ $p(D = \text{Yes} | H = \text{No}) = \frac{2}{7}$

▶ $p(S = \text{Yes} | H = \text{No}) = \frac{2}{7}$

Diabetes (D)	Smoke (S)	Heart Disease (H)
Y	N	Y
Y	N	N
N	Y	N
N	Y	N
N	N	N
N	Y	Y
N	N	N
N	Y	Y
N	N	N
Y	N	N

▶ Decision on $\mathbf{x} = [\text{Yes}, \text{Yes}]$ (a person that has diabetes and also smokes):

▶ $p(H = \text{Yes} | \mathbf{x}) \propto p(H = \text{Yes})p(D = \text{yes} | H = \text{Yes})p(S = \text{yes} | H = \text{Yes}) = 0.066$

▶ $p(H = \text{No} | \mathbf{x}) \propto p(H = \text{No})p(D = \text{yes} | H = \text{No})p(S = \text{yes} | H = \text{No}) = 0.057$

▶ Thus decide $H = \text{yes}$

Probabilistic classifiers

- ▶ How can we find the probabilities required in the Bayes decision rule?
- ▶ Probabilistic classification approaches can be divided in two main categories:
 - ▶ **Generative**
 - ▶ Estimate pdf $p(\mathbf{x}, \mathcal{C}_k)$ for each class \mathcal{C}_k and then use it to find $p(\mathcal{C}_k|\mathbf{x})$
 - or alternatively estimate both pdf $p(\mathbf{x}|\mathcal{C}_k)$ and $p(\mathcal{C}_k)$ to find $p(\mathcal{C}_k|\mathbf{x})$
 - ▶ **Discriminative**
 - ▶ Directly estimate $p(\mathcal{C}_k|\mathbf{x})$ for each class \mathcal{C}_k

Generative approach

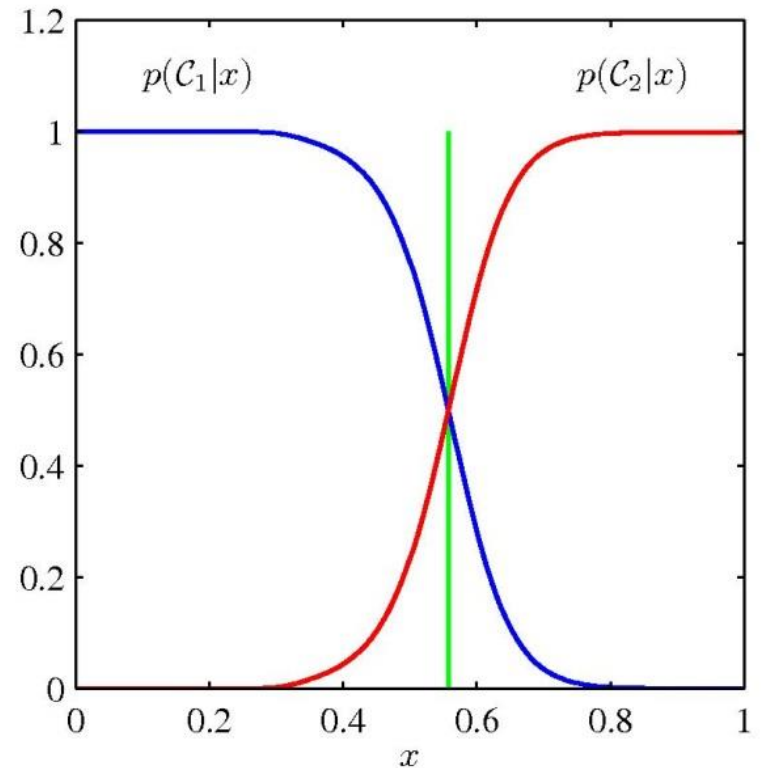
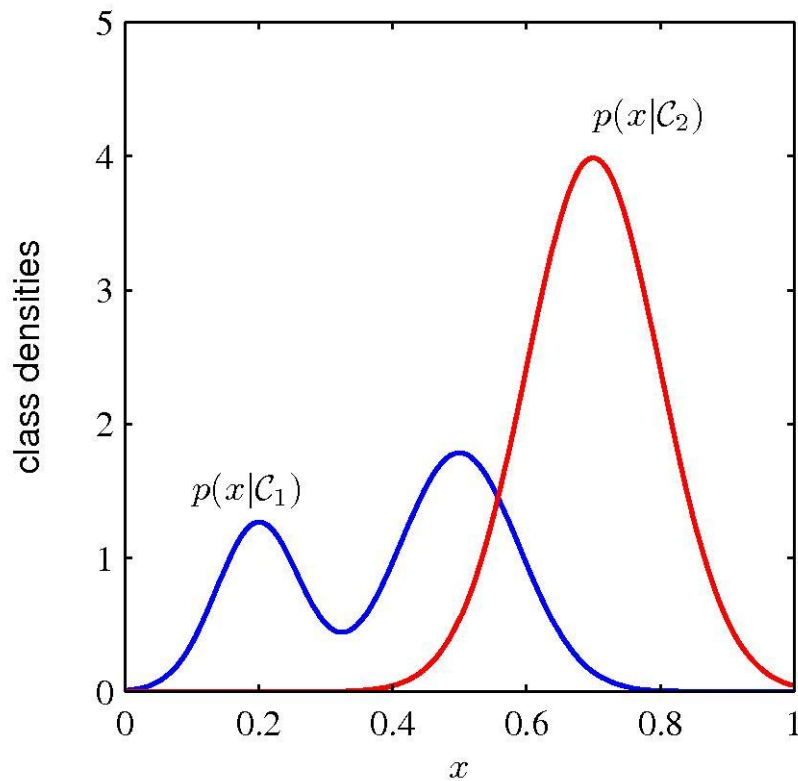
- ▶ Inference stage

- ▶ Determine class conditional densities $p(\mathbf{x}|\mathcal{C}_k)$ and priors $p(\mathcal{C}_k)$
- ▶ Use the Bayes theorem to find $p(\mathcal{C}_k|\mathbf{x})$

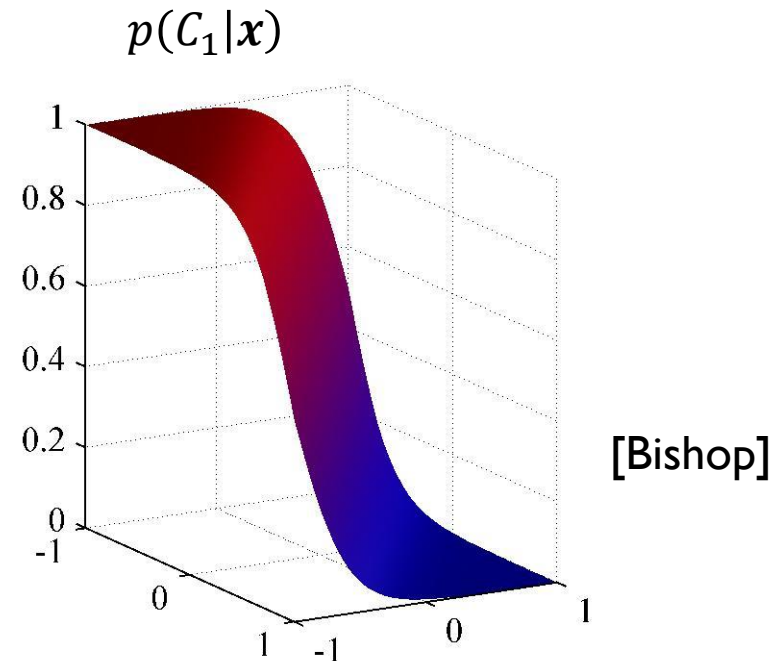
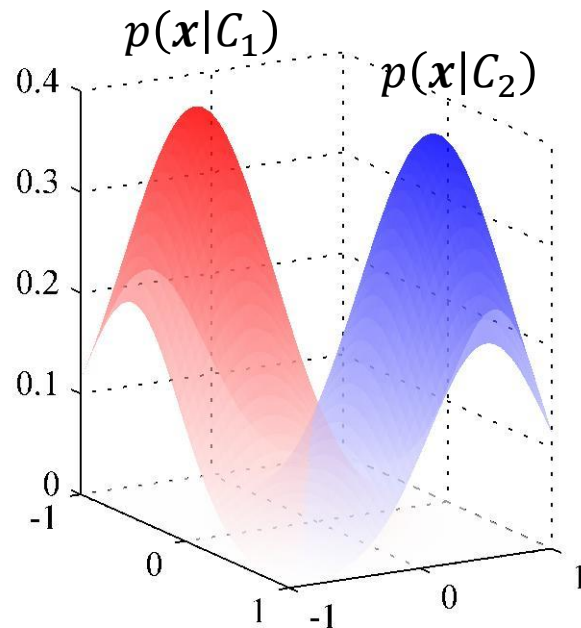
- ▶ Decision stage: After learning the model (inference stage), make optimal class assignment for new input

- ▶ if $p(\mathcal{C}_i|\mathbf{x}) > p(\mathcal{C}_j|\mathbf{x}) \quad \forall j \neq i$ then decide \mathcal{C}_i

Discriminative vs. generative approach



Class conditional densities vs. posterior



$$p(\mathcal{C}_1|\mathbf{x}) = \sigma(\mathbf{w}^T \mathbf{x} + w_0)$$

$$\mathbf{w} = \boldsymbol{\Sigma}^{-1}(\boldsymbol{\mu}_1 - \boldsymbol{\mu}_2)$$

$$w_0 = -\frac{1}{2}\boldsymbol{\mu}_1^T \boldsymbol{\Sigma}^{-1} \boldsymbol{\mu}_1 + \frac{1}{2}\boldsymbol{\mu}_2^T \boldsymbol{\Sigma}^{-1} \boldsymbol{\mu}_2 + \ln \frac{p(\mathcal{C}_1)}{p(\mathcal{C}_2)}$$

Discriminative approach

- ▶ Inference stage
 - ▶ Determine the posterior class probabilities $P(\mathcal{C}_k|\mathbf{x})$ directly
- ▶ Decision stage: After learning the model (inference stage), make optimal class assignment for new input
 - ▶ if $P(\mathcal{C}_i|\mathbf{x}) > P(\mathcal{C}_j|\mathbf{x}) \quad \forall j \neq i$ then decide \mathcal{C}_i

Posterior probabilities

- ▶ Two-class: $p(\mathcal{C}_k|\mathbf{x})$ can be written as a logistic sigmoid for a wide choice of $p(\mathbf{x}|\mathcal{C}_k)$ distributions

$$p(\mathcal{C}_1|\mathbf{x}) = \sigma(a(\mathbf{x})) = \frac{1}{1 + \exp(-a(\mathbf{x}))}$$

- ▶ Multi-class: $p(\mathcal{C}_k|\mathbf{x})$ can be written as a soft-max for a wide choice of $p(\mathbf{x}|\mathcal{C}_k)$

$$p(\mathcal{C}_k|\mathbf{x}) = \frac{\exp(a_k(\mathbf{x}))}{\sum_{j=1}^K \exp(a_j(\mathbf{x}))}$$

Discriminative approach: logistic regression

- ▶ More general than discriminant functions:

$K = 2$

- ▶ $f(\mathbf{x}; \mathbf{w})$ predicts posterior probabilities $P(y = 1|\mathbf{x})$

$$f(\mathbf{x}; \mathbf{w}) = \sigma(\mathbf{w}^T \mathbf{x})$$

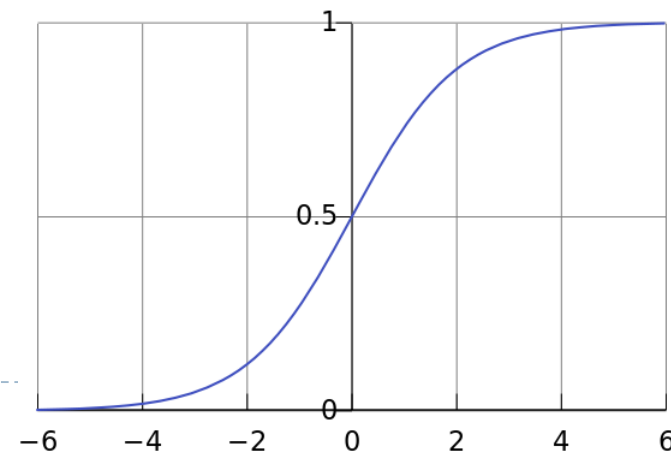
$$\mathbf{x} = [1, x_1, \dots, x_d]$$
$$\mathbf{w} = [w_0, w_1, \dots, w_d]$$

$\sigma(\cdot)$ is an activation function

- ▶ Sigmoid (logistic) function

- ▶ Activation function

$$\sigma(z) = \frac{1}{1 + e^{-z}}$$



Logistic regression

- ▶ $f(\mathbf{x}; \mathbf{w})$: probability that $y = 1$ given \mathbf{x} (parameterized by \mathbf{w})

$$P(y = 1 | \mathbf{x}, \mathbf{w}) = f(\mathbf{x}; \mathbf{w})$$

$K = 2$
 $y \in \{0, 1\}$

$$P(y = 0 | \mathbf{x}, \mathbf{w}) = 1 - f(\mathbf{x}; \mathbf{w})$$

$$f(\mathbf{x}; \mathbf{w}) = \sigma(\mathbf{w}^T \mathbf{x})$$

$$0 \leq f(\mathbf{x}; \mathbf{w}) \leq 1$$

estimated probability of $y = 1$ on input \mathbf{x}

- ▶ Example: Cancer (Malignant, Benign)
 - ▶ $f(\mathbf{x}; \mathbf{w}) = 0.7$
 - ▶ 70% chance of tumor being malignant

Logistic regression: Decision surface

- ▶ Decision surface $f(\mathbf{x}; \mathbf{w}) = \text{constant}$
 - ▶ $f(\mathbf{x}; \mathbf{w}) = \sigma(\mathbf{w}^T \mathbf{x}) = \frac{1}{1+e^{-(\mathbf{w}^T \mathbf{x})}} = 0.5$
- ▶ Decision surfaces are linear functions of \mathbf{x}

if $f(\mathbf{x}; \mathbf{w}) \geq 0.5$ then $y = 1$
else $y = 0$

Equivalent to

if $\mathbf{w}^T \mathbf{x} + w_0 \geq 0$ then $y = 1$
else $y = 0$

Logistic regression: ML estimation

- ▶ Maximum (conditional) log likelihood:

$$\hat{\mathbf{w}} = \operatorname{argmax}_{\mathbf{w}} \log \prod_{i=1}^n p(y^{(i)} | \mathbf{w}, \mathbf{x}^{(i)})$$

$$p(y^{(i)} | \mathbf{w}, \mathbf{x}^{(i)}) = f(\mathbf{x}^{(i)}; \mathbf{w})^{y^{(i)}} (1 - f(\mathbf{x}^{(i)}; \mathbf{w}))^{(1-y^{(i)})}$$

$$\begin{aligned} & \log p(\mathbf{y} | \mathbf{X}, \mathbf{w}) \\ &= \sum_{i=1}^n \left[y^{(i)} \log \left(f(\mathbf{x}^{(i)}; \mathbf{w}) \right) + (1 - y^{(i)}) \log \left(1 - f(\mathbf{x}^{(i)}; \mathbf{w}) \right) \right] \end{aligned}$$

Logistic regression: cost function

$$\hat{\mathbf{w}} = \underset{\mathbf{w}}{\operatorname{argmin}} J(\mathbf{w})$$

$$\begin{aligned} J(\mathbf{w}) &= - \sum_{i=1}^n \log p(y^{(i)} | \mathbf{w}, \mathbf{x}^{(i)}) \\ &= \sum_{i=1}^n -y^{(i)} \log \left(f(\mathbf{x}^{(i)}; \mathbf{w}) \right) - (1 - y^{(i)}) \log \left(1 - f(\mathbf{x}^{(i)}; \mathbf{w}) \right) \end{aligned}$$

- ▶ No closed form solution for

$$\nabla_{\mathbf{w}} J(\mathbf{w}) = 0$$

- ▶ However $J(\mathbf{w})$ is convex.

Logistic regression: Gradient descent

$$\mathbf{w}^{t+1} = \mathbf{w}^t - \eta \nabla_{\mathbf{w}} J(\mathbf{w}^t)$$

$$\nabla_{\mathbf{w}} J(\mathbf{w}) = \sum_{i=1}^n (f(\mathbf{x}^{(i)}; \mathbf{w}) - y^{(i)}) \mathbf{x}^{(i)}$$

- Is it similar to gradient of SSE for linear regression?

$$\nabla_{\mathbf{w}} J(\mathbf{w}) = \sum_{i=1}^n (\mathbf{w}^T \mathbf{x}^{(i)} - y^{(i)}) \mathbf{x}^{(i)}$$

Logistic regression: loss function

$$\text{Loss}(y, f(\mathbf{x}; \mathbf{w})) = -y \times \log(f(\mathbf{x}; \mathbf{w})) - (1 - y) \times \log(1 - f(\mathbf{x}; \mathbf{w}))$$

$$\text{Since } y = 1 \text{ or } y = 0 \Rightarrow \text{Loss}(y, f(\mathbf{x}; \mathbf{w})) = \begin{cases} -\log(f(\mathbf{x}; \mathbf{w})) & \text{if } y = 1 \\ -\log(1 - f(\mathbf{x}; \mathbf{w})) & \text{if } y = 0 \end{cases}$$

How is it related to zero-one loss?

$$\text{Loss}(y, \hat{y}) = \begin{cases} 1 & y \neq \hat{y} \\ 0 & y = \hat{y} \end{cases}$$

$$f(\mathbf{x}; \mathbf{w}) = \frac{1}{1 + \exp(-\mathbf{w}^T \mathbf{x})}$$

Logistic regression: cost function (summary)

- ▶ Logistic Regression (LR) has a more proper cost function for classification than SSE and Perceptron
- ▶ Why is the cost function of LR also more suitable than?

$$J(\mathbf{w}) = \frac{1}{n} \sum_{i=1}^n \left(y^{(i)} - f(\mathbf{x}^{(i)}; \mathbf{w}) \right)^2$$

where $f(\mathbf{x}; \mathbf{w}) = \sigma(\mathbf{w}^T \mathbf{x})$

- ▶ The conditional distribution $p(y|\mathbf{x}, \mathbf{w})$ in the classification problem is not Gaussian (it is Bernoulli)
- ▶ The cost function of LR is also convex

Multi-class logistic regression

- ▶ For each class k , $f_k(\mathbf{x}; \mathbf{W})$ predicts the probability of $y = k$
 - ▶ i.e., $P(y = k | \mathbf{x}, \mathbf{W})$
- ▶ On a new input \mathbf{x} , to make a prediction, pick the class that maximizes $f_k(\mathbf{x}; \mathbf{W})$:

$$\alpha(\mathbf{x}) = \operatorname{argmax}_{k=1,\dots,K} f_k(\mathbf{x})$$

if $f_k(\mathbf{x}) > f_j(\mathbf{x}) \quad \forall j \neq k$ then
decide C_k

Multi-class logistic regression

$$K > 2$$

$$y \in \{1, 2, \dots, K\}$$

$$f_k(\mathbf{x}; \mathbf{W}) = p(y = k | \mathbf{x}) = \frac{\exp(\mathbf{w}_k^T \mathbf{x})}{\sum_{j=1}^K \exp(\mathbf{w}_j^T \mathbf{x})}$$

- ▶ Normalized exponential (aka softmax)

- ▶ If $\mathbf{w}_k^T \mathbf{x} \gg \mathbf{w}_j^T \mathbf{x}$ for all $j \neq k$ then $p(C_k | \mathbf{x}) \simeq 1, p(C_j | \mathbf{x}) \simeq 0$

$$p(C_k | \mathbf{x}) = \frac{p(\mathbf{x} | C_k) p(C_k)}{\sum_{j=1}^K p(\mathbf{x} | C_j) p(C_j)}$$

Logistic regression: multi-class

$$\widehat{\mathbf{W}} = \operatorname{argmin}_{\mathbf{W}} J(\mathbf{W})$$

$$\begin{aligned} J(\mathbf{W}) &= -\log \prod_{i=1}^n p(\mathbf{y}^{(i)} | \mathbf{x}^{(i)}, \mathbf{W}) \\ &= -\log \prod_{i=1}^n \prod_{k=1}^K f_k(\mathbf{x}^{(i)}; \mathbf{W})^{y_k^{(i)}} \\ &= -\sum_{i=1}^n \sum_{k=1}^K y_k^{(i)} \log(f_k(\mathbf{x}^{(i)}; \mathbf{W})) \end{aligned}$$

\mathbf{y} is a vector of length K (1-of- K coding)
e.g., $\mathbf{y} = [0, 0, 1, 0]^T$ when the target class is C_3

$$\mathbf{W} = [\mathbf{w}_1 \quad \cdots \quad \mathbf{w}_K]$$

$$\mathbf{Y} = \begin{bmatrix} \mathbf{y}^{(1)} \\ \vdots \\ \mathbf{y}^{(n)} \end{bmatrix} = \begin{bmatrix} y_1^{(1)} & \cdots & y_K^{(1)} \\ \vdots & \ddots & \vdots \\ y_1^{(n)} & \cdots & y_K^{(n)} \end{bmatrix}$$

Logistic regression: multi-class

$$\mathbf{w}_j^{t+1} = \mathbf{w}_j^t - \eta \nabla_{\mathbf{W}} J(\mathbf{W}^t)$$

$$\nabla_{\mathbf{w}_j} J(\mathbf{W}) = \sum_{i=1}^n \left(f_j(\mathbf{x}^{(i)}; \mathbf{W}) - y_j^{(i)} \right) \mathbf{x}^{(i)}$$

Logistic Regression (LR): summary

- ▶ LR is a linear classifier
- ▶ LR optimization problem is obtained by maximum likelihood
 - ▶ when assuming Bernoulli distribution for conditional probabilities whose mean is $\frac{1}{1+e^{-(w^T x)}}$
- ▶ No closed-form solution for its optimization problem
 - ▶ But convex cost function and global optimum can be found by gradient ascent

Discriminative vs. generative: number of parameters

- ▶ d -dimensional feature space
- ▶ Logistic regression: $d + 1$ parameters
 - ▶ $\mathbf{w} = (w_0, w_1, \dots, w_d)$
- ▶ Generative approach:
 - ▶ Gaussian class-conditionals with shared covariance matrix
 - ▶ $2d$ parameters for means
 - ▶ $d(d + 1)/2$ parameters for shared covariance matrix
 - ▶ one parameter for class prior $p(C_1)$.
- ▶ But LR is more robust, less sensitive to incorrect modeling assumptions

Summary of alternatives

▶ Generative

- ▶ Most demanding, because it finds the joint distribution $p(\mathbf{x}, \mathcal{C}_k)$
- ▶ Usually needs a large training set to find $p(\mathbf{x}|\mathcal{C}_k)$
- ▶ Can find $p(\mathbf{x}) \Rightarrow$ Outlier or novelty detection

▶ Discriminative

- ▶ Specifies what is really needed (i.e., $p(\mathcal{C}_k|\mathbf{x})$)
- ▶ More computationally efficient

Resources

- ▶ C. Bishop, “Pattern Recognition and Machine Learning”, Chapter 4.2-4.3.